POLS/CS&SS 503: Advanced Quantitative Political Methodology

LINEAR REGRESSION ESTIMATOR

April 14, 2015

Jeffrey B. Arnold



CENTER for **STATISTICS** and the **SOCIAL SCIENCES**



Overview

Regression Coefficient Anatomy

Linear Regression Population Model

Sampling Distribution

Estimators and mean squared error

Gauss-Markov Theorem

Regression Coefficient Anatomy

Linear Regression Population Model

Sampling Distribution

Estimators and mean squared error

Gauss-Markov Theorem

Coeficients of a simple regression

$$Y_i = A + BX_i + E_i$$

The least squares coefficients are

$$A = \bar{Y} + B\bar{X}$$
$$B = \frac{\sum_{i} (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i} (X_i - \bar{X})^2}$$

- State *B* in terms of covariance of *X* and *Y* and variance?
- State B in terms of correlation of X and Y and standard deviations?
- What values can B take if SD(X) = SD(Y) = 1?
- What is \hat{Y} for $X = \bar{X}$?
- What happens to B as $\mathrm{V}(X)$ decreases? $\mathrm{V}(Y)$ decreases? If $\mathrm{V}(X)=0$

Least squares when VX = 0



Least squares coefficients are unidentified if V x = 0

- If V x = 0 then least squares solution is unidentified
- There is no unique value of A,B that $\arg\min_{A,B}\sum_i E_i^2$

```
y <- c(1, 2, 3, 4, 5)
x <- 1
ybar <- mean(y)
ybar
## [1] 3
# A = 2, B = 1
sum((y - 2 - 1 * x) ^ 2)
## [1] 10
# A = -7, B = 10
sum((y + 7 - 10 * x) ^ 2)
## [1] 10
```

Coefficients of a multiple regression

$$\vec{Y} = Xb + e$$

- + $oldsymbol{b} = \left(oldsymbol{X}'oldsymbol{X}
 ight)^{-1}oldsymbol{X}'oldsymbol{y}$. Not that intuitive!
- Coefficient $oldsymbol{b}_j$ is

$$oldsymbol{b}_k = rac{\mathsf{C}(oldsymbol{y}, ilde{oldsymbol{x}}_j)}{\mathsf{V}(ilde{oldsymbol{x}}_k)}$$

• Where $ilde{m{x}}_j$ are the residulals of $m{x}_j$ on all X_h where h
eq j

$$\tilde{X}_{j,i} = X_{j,i} - \tilde{A} - \sum_{h \neq j} \tilde{B}_h X_h$$

Regression example

See multiple_regression_anatomy.R

Least Squares coefficients are unidentified if $\left(X'X\right)^{-1}$ does not exist

- Common cases in which $(X'X)^{-1}$ does not exist:
 - Number of observations less than k+1
 - X_k is constant
 - X_k is a linear function of other variables: $X_k = \sum_{j \neq k} c_j X_j$.
 - · dummy variables for all categories of a categorical variable
 - variable multiplied by the constant of another variable

Which of these would be cases of collinearity and why?

- There is a variable that takes values "white", "black", "hispanic", "asian", "other". You include a dummy variable for each category.
- GDP, GDP per capita, and population
- Log GDP, log GDP per capita, and log population
- GDP in millions of dollars; GDP in trillions of dollars
- GDP measured in nominal value; GDP measured in real terms
- · Regression with 3 variables and 4 observations

Regression Coefficient Anatomy

Linear Regression Population Model

Sampling Distribution

Estimators and mean squared error

Gauss-Markov Theorem

Population model in a simple regression

$$Y_i = \alpha + \beta X_i + \epsilon_i$$

Assumptions for statistical inference

- 1. X is not invariant: V(X) > 0
- 2. *Linearity*. Average value of error given x is 0. $E(\epsilon_i) = E(\epsilon_i | x_i) = 0$

$$\mu_i = E(Y_i) = \mathsf{E}(Y|X_i) = \mathsf{E}(\alpha + \beta X_i + \epsilon_i) = \alpha + \beta x_i$$

3. Constant variance $\mbox{Variance}$ of the errors is the same regardless of the value of X

$$\mathsf{V}(Y|x_i) = E(\epsilon_i^2) = \sigma_\epsilon^2$$

- 4. Independence: Observations are sampled independently. $Cor(\epsilon_i, \epsilon_j) = 0$ for all $i \neq j$.
- 5. Fixed X or X measured without error and independent of the error.
- 6. Errors are normally distributed $\epsilon_i \sim N\left(0, \sigma_{\epsilon}^2\right)$

Population model in a multiple regression

$$Y_i = \alpha + \beta_1 x_{i,1} + \beta_2 x_{i,2} + \dots + \beta_k x_{i,k} + \epsilon_i$$

Assumptions for statistical inference

- 1. *X* is not invariant and no X is a perfect linear function of the others.
- 2. Linearity. $E(\epsilon_i) = 0$
- 3. Constant variance $V(\epsilon_i) = \sigma_{\epsilon}^2$
- 4. Independence Observations are sampled independently. $Cor(\epsilon_i, \epsilon_j) = 0$ for all $i \neq j$.
- 5. Fixed X or X measured without error and independent of the error
- 6. *Normality* Errors are normally distributed $\epsilon_i \sim N\left(0, \sigma_{\epsilon}^2\right)$

Regression Coefficient Anatomy

Linear Regression Population Model

Sampling Distribution

Estimators and mean squared error

Gauss-Markov Theorem

Definitions

population The observations of interest. May be theoretical sample The data you have.
 parameter A function of the population distribution statistic A function of the sample
 sampling distribution The distribution of a statistic calculated from the distribution of samples of a given size drawn from a population.

See Sampling_Distributions.Rmd

Sampling Distribution of Simple Regression Coefficients

The sampling distributions of A, B given $Y_i = \alpha + \beta X_i + \epsilon_i$

expected values (linearity)

$$\begin{split} \mathbf{E}(A) &= \alpha \\ \mathbf{E}(B) &= \beta \end{split}$$

· variances (linearity, constant variance, independence)

$$V(A) = \frac{\sigma_{\epsilon}^2}{n} \cdot \frac{\sum x_i^2}{\sum (x_i - \bar{x})^2}$$
$$V(B) = \frac{\sigma_{\epsilon}^2}{\sum (x_i - \bar{x})^2} = \frac{\sigma_{\epsilon}^2}{n \, V(x)}$$

normal distribution (normal errors)

 $\begin{aligned} A &\sim N(\mathsf{E}(A), \mathsf{V}(A)) \\ B &\sim N(\mathsf{E}(B), \mathsf{V}(B)) \end{aligned}$

Coefficient sampling distributions in multiple regression

The sampling distributions of B_k given $Y_i = \alpha + \sum \beta_j X_{j,i} + \epsilon_i$

- Expected value: $E(B_K) = \beta_k$
- Variance:

$$V(B_j) = \frac{1}{(1 - R_j^2)} \frac{\sigma_\epsilon^2}{\sum (x_{i,j} - \bar{x}_j)^2}$$
$$= \frac{\sigma_\epsilon^2}{\sum_i (x_{i,j} - \hat{x}_{i,j})^2}$$

Where R_j^2 is R^2 from regression of X_j on other X, and \hat{x}_{ij} are fitted values from that regression.

- Normally distributed if errors are normally distributed or as $n \to \infty$.
- $\cdot ~ m{b}$ is multivariate normally distributed

$$\boldsymbol{b} \sim N\left(\boldsymbol{\beta}, \sigma_{\epsilon}^2 (X'X)^{-1}\right)$$

Regression Coefficient Anatomy

Linear Regression Population Model

Sampling Distribution

Estimators and mean squared error

Gauss-Markov Theorem

Let's define some things

statistic Function of a sample, e.g. Sample mean $ar{x}=rac{1}{n}\sum x_i$

- parameter Function of the population distribution, e.g. Expected value μ of the normal distribution.
- estimator Method to use a sample statistic (estimate) to infer a population parameter (estimand)

How to determine if an estimator is good?

- Is $\hat{eta} = \left(oldsymbol{X}'oldsymbol{X}
 ight)^{-1}oldsymbol{X}'oldsymbol{y}$ a good estimator for eta?
- · Would another estimator be better?
- First, need criteria to by which to judge estimators

What makes an estimator good?

- Bias
- Variance
- Efficiency (mean squared error)
- Consistency

Bias and Variance

Bias

On average how far off is the estimator?

$$\mathsf{bias}(\hat{eta}) = \mathsf{E}(\hat{eta}) - eta$$

Variance

Does the estimator give similar results in different samples?

$$\mathsf{V}(\hat{\beta}) = \mathsf{E}\left(\left(\beta - \mathsf{E}(\hat{\beta})\right)^2\right)$$

Bias and Variance Visualized



Bias

What makes an estimator good?

- Unbiased methods may still miss the truth by a large amount, just direction not systematic
- Unbiased estimates can be horrible: random draw from numbers 0–24 for time of day
- Biased estimates are not necessarily terrible: a clock that's 2 minutes fast

You may prefer a biased, low variance estimator to an unbiased, high variance estimator



Mean Squared Error (MSE)

MSE is

$$MSE(\hat{\beta}) = \mathsf{E}\left((\hat{\beta} - \beta)^2\right)$$

MSE trades off bias and variance

$$\begin{split} MSE(\hat{\beta}) &= \mathsf{E}((\hat{\beta} - \mathsf{E}(\hat{\beta}))^2) + \mathsf{E}(\mathsf{E}(\hat{\beta}) - \beta))^2 \\ &= \mathsf{V}(\hat{\beta}) + \left(\mathsf{bias}(\hat{\beta}, \beta)\right)^2 \end{split}$$

- root mean squared error (RMSE) $\sqrt{\rm MSE}$: on average how far is an estimate from the truth
- · An efficient estimator has the smallest MSE
- What is the MSE of an unbiased estimator?

$$MSE(\hat{\beta}) = \mathsf{V}(\hat{\beta}) + \left(\mathsf{bias}(\hat{\beta},\beta)\right)^2 = \mathsf{V}(\hat{\beta}) + 0 = \mathsf{V}(\hat{\beta})$$

MSE Example

- Suppose population parameter $\beta=1$
- Consider two estimators $\hat{\beta}_1$ and $\hat{\beta}_2$.

•
$$\hat{\beta}_1 \sim N(1, 1^2)$$

- $\hat{\beta}_2 \sim N(0.5, 0.5^2)$
- What are the bias, variance, and MSE of each estimator?

Consistency

• A consistent estimator converges to the parameter value as the number of observations grows

$$\mathsf{E}(\hat{\beta}-\beta)\to 0 \qquad \qquad \text{as } n\to\infty$$

- A concern of econometricians
- May not be as much a concern in finite, small sample sizes
- We will mainly be concerned with efficiency, secondarily with bias, rarely with consistency

Regression Coefficient Anatomy

Linear Regression Population Model

Sampling Distribution

Estimators and mean squared error

Gauss-Markov Theorem

LS assumptions and consequences of violations

	Assumption		Consequence of violation
1	No perfect collinearity	$rank(oldsymbol{X}) = k$, $k < n$	Coefficients unidentified
2	$oldsymbol{X}$ is exogenous	$E(\boldsymbol{X}\epsilon) = 0$	Biased, even as $n o \infty$
3	Disturbances have mean 0	$E(\epsilon) = 0$	Biased, even as $n o \infty$
4	No serial correlation	$E(\epsilon_i \epsilon_j) = 0, i \neq j$	Unbiased but ineff. Wrong se.
5	Homoskedastic errors	$E(\epsilon'\epsilon') = \sigma^2 I$	Unbiased but ineff. Wrong se.
6	Normal errors	$\epsilon \sim N(0, \sigma^2)$	se wrong unless $n \to \infty$

Assumptions stronger from top to bottom, 4 and 5 could be combined

Unbiasedness of LS

- Only need assumptions 1-3 (no collinearity, $oldsymbol{X}$ exogenous, $\mathsf{E}(\epsilon)=0$
- Start with

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'(\mathbf{X}\beta + \epsilon)$$
$$= \beta + (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\epsilon$$

Take the expectation

$$\begin{split} \mathsf{E}(\hat{\beta}) &= \mathsf{E}(\beta) + \mathsf{E}(\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{\epsilon}) \\ &= \mathsf{E}(\beta) + \boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\,\mathsf{E}(\boldsymbol{\epsilon}) \\ &= \mathsf{E}(\beta) \end{split}$$

• Since E
$$(\hat{eta})=$$
 E (eta) , LS is unbiased.

Gauss-Markov

- If make assumptions 1–5: LS is the best linear unbiased estimator (BLUE)
- + LS estimator is **linear** because $\hat{eta} = M y$, where $M = \left(X' X
 ight)^{-1} X'$
- **best** is best mean squared error (MSE).
- If LS is unbiased, then its mean squared error is the same as its ...?
- Could exist other non-linear unbiased estimators with smaller MSE, e.g. Robust regression when population has fat tailed errors
- If errors are Gaussian, LS is Minimum Variance Unbiased (MVU).
- MVU = for *all* estimators that are unbiased. $\hat{\beta}$ has smallest variance (and MSE).

References

- Some slides derived from Christopher Adolph *Linear Regression in Matrix Form / Propoerties & Assumptions of Linear Regression*. Used with permission.
 - <http://faculty.washington.edu/cadolph/503/topic3.pw.pdf>
- Material included from
 - Fox Ch 6, 9.3
 - Angrist and Pischke, Chapter 3

dasfjasda